Zig-Zag Sampler A MCMC Game-Changer

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Today's Menu

- I The Zig-Zag Sampler: What Is It?
- 2 The Algorithm: How to Use It?
- 3 Proof of Concept: How Good Is It?

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I The Zig-Zag Sampler: What Is It?

A continuous-time variant of MCMC algorithms

Zig-Zag

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- I.I Keywords: PDMP (1/2) <u>PDMP</u> (Piecewise Deterministic¹ Markov Process²) (Davis, 1984)
- I. Mostly deterministic with the exception of random jumps happens at random times
- 2. Continuous-time, instead of discrete-time processes
- → Plays a complementary role to SDEs / Diffusions

Property	PDMP	SDE
Exactly simulatable?	\checkmark	×
Subject to discretization errors?	×	\checkmark
Driving noise Hirofumi Shiba	Poisson	Gauss

i History of PDMP Applications

- I. First applications: control theory, operations research, etc. (Davis, 1993)
- 2. Second applications: Monte Carlo simulation in material sciences (Peters and de With, 2012)
- 3. Third applications: Bayesian statistics (Bouchard-Côté et al., 2018)
- I. Mostly deterministic with the exception of random jumps happens at random times
- 2. Continuous-time, instead of discrete-time processes

I.2 Keywords: PDMP (2/2)

- We will concentrate on Zig-Zag sampler (Bierkens, Fearnhead, et al., 2019)
- Other PDMPs: Bouncy sampler (Bouchard-Côté et al., 2018), Boomerang sampler (Bierkens et al., 2020)



The most famous three PDMPs. Animated by (Grazzi, 2020)

I.3 Menu

What We've Learned

The new algorithm 'Zig-Zag Sampler' is based on comtinuous-time process called PDMP.

What We'll Learn in the Rest of this Section

We will review 3 instances of the standard (discrete-time) MCMC algorithm: MH, Lifted MH, and MALA.

- I. Review: MH (Metropolis-Hastings) algorithm
- 2. Review: Lifted MH, A method bridging MH and Zig-Zag
- 3. Comparison: MH vs. Lifted MH vs. Zig-Zag
- 4. Review: MALA (Metropolis Adjusted Langevin Algorithm)
- 5. Comparison: Zig-Zag vs. MALA

I.4 Review: Metropolis-Hastings (1/2)

(Metropolis et al., 1953)-(Hastings, 1970)

Input: Target distribution p, (symmetric) proposal distribution q

I. Draw a $X_t \sim q(-|X_{t-1})$

2. Compute

$$lpha(X_{t-1},X_t)=rac{p(X_t)}{p(X_{t-1})},$$

- 3. Draw a uniform random number $U \sim U([0,1])$.
- 4. If $lpha(X_{t-1},X_t) \leq U$, then $X_t \leftarrow X_{t-1}$. Do nothing otherwise.

5. Return to Step 1.

MH algorithm works even without p's normalizing constant. Hence, its ubiquity.

I.5 Review: Metropolis-Hastings (2/2)

Alternative View: MH is a generic procedure to turn a simple q-Markov chain into a Markov chain converging to p.

The Choise of Proposal q

• <u>Random Walk Metropolis</u> (Metropolis et al., 1953): Uniform / Gaussian

$$q(y|x) = q(y-x) \in \left\{rac{d\mathrm{U}([0,1])}{d\lambda}(y-x), rac{d\mathrm{N}(0,\Sigma)}{d\lambda}(y-x)
ight\}$$

• Hybrid / Hamiltonian Monte Carlo (Duane et al., 1987): Hamiltonian dynamics

 $q(y|x)=\delta_{x+\epsilon
ho},\qquad \epsilon>0,\
ho: ext{momentum defined via Hamiltonian}$

• <u>Metropolis-adjusted Langevin algorithm</u> (MALA) (Besag, 1994): Langevin diffusion

 $q(-|X_t):= ext{ the transition probability of } X_t ext{ where } dX_t=
abla \log p(X_t)\,dt+\sqrt{2eta^{-1}}dB_t.$

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I.6 Problem: Reversibility

Reversibility (a.k.a detailed balance):

$$p(x)q(x|y) = p(y)q(y|x).$$

In words:

 $\operatorname{Probability}[\operatorname{Going} x \to y] = \operatorname{Probability}[\operatorname{Going} y \to x].$

 \rightarrow Harder to explore the entire space

 \rightarrow Slow mixing of MH

From the beginning of 21th century, many-efforts have been made to make MH irreversible

I.7 Lifting (I/3)

Lifting: A method to make MH's dynamics irreversible

How?: By adding an auxiliary variable $\sigma \in \{\pm 1\}$, called momentum

Lifted MH (Turitsyn et al., 2011)

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Input: Target p, two proposals q^{(+1)}, q^{(-1)}, and momentum \sigma \in \{\pm 1\}
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- I. Draw X_t from $q^{(\sigma)}$
- 2. Do a MH step
- 3. If accepted, go back to Step 1.
- 4. If rejected, flip the momentum and go back to Step 1.



I.8 Lifting (2/3)

 $q^{(+1)}$: Only propose ightarrow moves $q^{(-1)}$: Only propose \leftarrow moves

 \rightarrow Once going uphill, it continues to go uphill.

 \rightarrow This is irreversible, since

 $\operatorname{Probability}[x
ightarrow y]$ \neq Probability $[y \rightarrow x]$.

Reversible dynamic of MH has 'irreversified'



Lifted MH successfully explores the edges of the

*Irreversibility actually improves the efficiency of MCMC, as we observe in two slides later. \Im

1.10 Comparison: MH vs. LMH vs. Zig-Zag (1/2)



Zig-Zag corresponds to the limiting case of lifted MH as the step size of proposal q goes to zero, as we'll learn later.

→ Zig-Zag has a maximum irreversibility.

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I.II Comparison: MH vs. LMH vs. Zig-Zag (2/2) Irreversibility actually improves the efficiency of MCMC. Faster decay of **autocorrelation** $\rho_t \approx \operatorname{Corr}[X_0, X_t]$ implies

- I. faster mixing of MCMC
- 2. lower variance of Monte Carlo estimates



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MH

Lifted MH

Zig-Zag

I.I2 Review: MALA

Langevin diffusion: A diffusion process defined by the following SDE:

 $dX_t =
abla \log p(X_t) \, dt + \sqrt{2eta^{-1}} dB_t.$

Langevin diffusion itself converges to the target distribution p in the sense that ¹

 $\|p_t-p\|_{L^1} o 0, \qquad t o\infty.$

Two MCMC algorithms derived from Langevin diffusion:

<u>ULA (Unadjusted Langevin Algorithm)</u>

Use the discretization of (X_t) . Discretization errors accumulate.

MALA (Metropolis Adjusted Langevin Algorithm)

Use ULA as a proposal in MH, erasing the errors by MH steps.

I. under fairly general conditions on p.

I.I3 Comparison: Zig-Zag vs. MALA (I/3) How fast do they go back to high-probability regions? ^I



Zig-Zag

MALA

Irreversibility of Zig-Zag accelerates its convergence.

I. The target here is the standard Cauchy distribution C(0,1), equivalent to t(1) distribution. Its heavy tails hinder the convergence of MCMC.

I.I4 Comparison: Zig-Zag vs. MALA (2/3)



Caution: Fake Continuity The left plot looks continuous, but it actually is not.

MALA trajectory

MH, including MALA, is actually a <u>discrete-time process</u>. The plot is obtained by <u>connecting the points</u> by line segments.

I.I5 Comparison: Zig-Zag vs. MALA (3/3) Monte Carlo estimation is also done differently:



MALA outputs
$$(X_n)_{n\in[N]}$$
 defines

$$rac{1}{N}\sum_{n=1}^N f(X_n) \stackrel{N o \infty}{\longrightarrow} \int_{\mathbb{R}^d} f(x) p(x) \, dx.$$



$$\underline{Z}$$
ig- \underline{Z} ag outputs $(X_t)_{t\in[0,T]}$ defines $\int_0^T f(X_t)\,dt \stackrel{T o\infty}{\longrightarrow} \int_{\mathbb{R}^d} f(x)p(x)\,dx.$

I.I6 Recap of Section I

- Zig-Zag Sampler's trajectory is a <u>PDMP</u>.
- <u>PDMP</u>, by design, has maximum irreversibility.
- Irreversibility leads to faster convergence of Zig-Zag in comparisons against MH, Lifted MH, and especially MALA.



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MH

Lifted MH

Zig-Zag

2 The Algorithm: How to Use It?

Fast and exact simulation of continuous trajectory.

2.1 Review: MH vs. LMH vs. Zig-Zag (1/2)As we've learned before, Zig-Zag corresponds to the limitingcase of lifted MH as the step size of proposal q goes to zero.



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2.2 Review: MH vs. LMH vs. Zig-Zag (2/2) 'Limiting case of lifted MH' means that we only simulate <u>where</u> <u>we should flip the momentum</u> $\sigma \in \{\pm 1\}$ in Lifted MH.



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2.3 Algorithm (1/2)

'Limiting case of lifted MH' means that we only simulate <u>where</u> we should flip the momentum $\sigma \in \{\pm 1\}$ in Lifted MH.

(Id I Zig Zag sampler Bierkens, Fearnhead, et al., 2019)

Input: Gradient $\nabla \log p$ of log target density p

For $n\in\{1,2,\cdots,N\}$:

I. Simulate an first arrival time T_n of a Poisson point process (described in the next slide)

2. Linearly interpolate until time T_n :

$$X_t = X_{T_{n-1}} + \sigma(t - T_{n-1}), \qquad t \in [T_{n-1}, T_n].$$

3. Go back to Step I with the momentum $\sigma \in \{\pm 1\}$ flipped

I. Multidimensional extension is straightforward, but we won't cover it today. Hirofumi Shiba

2.4 Algorithm (2/2)

(Fundamental Property of Zig-Zag Sampler (Id) Bierkens, Fearnhead, et al., 2019)

Let $U(x) := -\log p(x)$. Simluating a Poisson point process with a rate function

$$\lambda(x,\sigma):= igg(\sigma U'(x)igg)_+ + \gamma(x)igg)$$

ensures the Zig-Zag sampler converges to the target p, where γ is an arbitrary non-negative function.

Its ergodicity is ensured as long as there exists c, C > 0 such that I

$$p(x) \leq C |x|^{-c}.$$

I. With some regularity conditions on U. (See Bierkens, Roberts, et al., 2019).

2.5 Core of the Algorithm

Given a rate function

$$\lambda(x,\sigma):=\!\left(\sigma U'(x)
ight)_+\!+\gamma(x)$$

how to simulate a corresponding Poisson point process?

What We'll Learn in the Rest of this Section 2

- I. What is **Poisson Point Process**?
- 2. How to Simulate It?
- 3. Core Technique: Poisson Thinning

Take Away: Zig-Zag sampling reduces to Poisson Thinning.

2.6 Simulating Poisson Point Process (1/2)

What is a Poisson Point Process with rate λ ?

The number of points in [0, t] follows a Poisson distribution with mean $\int_0^t \lambda(x_s, \sigma_s) \, ds$:

$$N([0,t]) \sim \mathrm{Pois}\left(M(t)
ight), \qquad M(t) := \int_0^t \lambda(x_s,\sigma_s)\,ds$$

We want to know when the first point T_1 falls on $[0,\infty)$.



When
$$\lambda(x,\sigma)\equiv c~(ext{constant})$$
,

- blue line: Poisson Process
- red dots: Poisson Point Process

satisfying
$$N_t = N([0,t]) \sim \mathrm{Pois}(ct).$$

2.7 Simulating Poisson Point Process (2/2)

Proposition (Simulation of Poisson Point Process)

The first arrival time T_1 of a Poisson Point Process with rate λ can be simulated by

$$T_1 \stackrel{ ext{d}}{=} M^{-1}(E), \qquad E \sim \operatorname{Exp}(1), M(t) := \int_0^t \lambda(x_s, \sigma_s) \, ds,$$

where Exp(1) denotes the exponential distribution with parameter 1.

Since
$$\lambda(x,\sigma):=\!\left(\sigma U'(x)
ight)_+\!+\gamma(x)$$
, M can be quite

complicated.

- \rightarrow Inverting M can be impossible.
- → We need more general techniques: Poisson Thinning.

2.8 Poisson Thinning (1/2)

(Lewis and Shedler, 1979)

To obtain the first arrival time T_1 of a Poisson Point Process with rate λ ,

I. Find a bound M that satisfies

$$m(t):=\int_0^t\lambda(x_s,\sigma_s)\,ds\leq M(t).$$

2. Simulate a point T from the Poisson Point Process with intensity M.

3. Accept T with probability $\frac{m(T)}{M(T)}$.

•
$$m(t)$$
: Defined via $\lambda(x,\sigma):= \left(\sigma U'(x)
ight)_+ + \gamma(x).$

• M(t): Simple upper bound $m \leq M$, such that M^{-1} is analytically tractable.

2.9 Poisson Thinning (2/2)

In order to simulate a Poisson Point Process with rate

$$\lambda(x,\sigma):= igg(\sigma U'(x)igg)_+ + \gamma(x),$$

we find a invertible upper bound M that satisfies

$$\int_0^t \lambda(x_s,\sigma_s)\,ds = m(t) \leq oldsymbol{M}(t).$$

for all possible Zig-Zag trajectories $\{(x_s, \sigma_s)\}_{s \in [0,T]}$.

2.10 Recap of Section 2

- I. Continuous-time MCMC, based on <u>PDMP</u>, has an entirely different algorithm and strategy.
- 2. To simulate <u>PDMP</u> is to simulate <u>Poisson Point Process</u>.
- 3. The core technology to simulate Poisson Point Process is Poisson Thinning.
- 4. Poisson Thinning is about finding an upper bound M, with tractable inverse M^{-1} ; Typically a polynomial function.
- 5. The upper bound M has to be given on a case-by-case basis.

3 Proof of Concept: How Good Is It?

Quick demonstration of the state-of-the-art performance on a toy example.

- **3.1 Review: The 3 Steps of Zig-Zag Sampling** Given a target *p*,
- I. Calculate the negative log-likelihood $U(x) := -\log p(x)$
- 2. Fix a refresh rate $\gamma(x)$ and compute the rate function

$$\lambda(x,\sigma):=\!\left(\sigma U'(x)
ight)_+\!+\gamma(x).$$

3. Find an invertible upper bound M that satisfies

$$\int_0^t \lambda(x_s,\sigma_s)\,ds =: m(t) \leq {M}(t).$$

3.2 Model: Id Gaussian Mean Reconstruction

Setting

• <u>Data</u>: $y_1, \cdots, y_n \in \mathbb{R}$ aquired by

 $y_i \stackrel{ ext{iid}}{\sim} \mathrm{N}(x_0, \sigma^2), \qquad i \in [n],$

with $\sigma>0$ known, $x_0\in\mathbb{R}$ unknown.

- <u>Prior</u>: $N(0, \rho^2)$ with known $\rho > 0$.
- <u>Goal</u>: Sampling from the posterior

$$p(x) \, \propto \, \left(\prod_{i=1}^n \phi(x|y_i,\sigma^2)
ight) \phi(x|0,
ho^2),$$

where $\phi(x|y,\sigma^2)$ is the ${
m N}(y,\sigma^2)$ density.

The negative log-likelihood: $U(x) = -\log p(x)$ $=rac{x^2}{2
ho^2}+rac{1}{2\sigma^2}\sum_{i=1}^n(x-y_i)^2$ $U'(x)=rac{x}{
ho^2}+rac{1}{\sigma^2}\sum_{i=1}^n(x-y_i),$ $U^{\prime\prime}(x)=rac{1}{
ho^2}+rac{n}{\sigma^2}.$

3.3 Menu

In the rest of this Section 3, we'll learn:

- I. Even a simple Zig-Zag Sampler with $\gamma\equiv 0$ surpasses MALA.
- 2. Incorporating sub-sampling, Zig-Zag with Control Variates further improves the efficiency.



3.4 Simple Zig-Zag Sampler with $\gamma \equiv 0$ (1/2) Fixing $\gamma \equiv 0$, we obtain the upper bound *M*

$$egin{aligned} m(t) &= \int_0^t \lambda(x_s,\sigma_s)\,ds = \int_0^t igg(\sigma U'(x_s)igg)_+\,ds \ &\leq igg(rac{\sigma x}{
ho^2} + rac{\sigma}{\sigma^2}\sum_{i=1}^n(x-y_i) + t\,igg(rac{1}{
ho^2} + rac{n}{\sigma^2}igg)igg)_+ \ &=: (a+bt)_+ = M(t), \end{aligned}$$

where

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3.5 Result: Id Gaussian Mean Reconstruction

We generated 100 samples from $N(x_0, \sigma^2)$ with $x_0 = 1$.



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3.6 MSE per Epoch: The Vertical Axis MSE (Mean Squared Error) of $\{X_i\}_{i=1}^n$ is defined as

$$rac{1}{n}\sum_{i=1}^n (X_i-x_0)^2.$$

Epoch: Unit computational cost.

The following is considered as one epoch:

• One evaluation of a likelihood ratio

 $\frac{p(X_{n+1})}{p(X_n)}.$

• One evaluation of a Poisson Point Process.

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3.7 Good News!

Case-by-case construction of an upper bound M is too complicated / demanding.

Therefore, we are trying to automate the whole procedure.

Automatic Zig-Zag

- I. Automatic Zig-Zag (Corbella et al., 2022)
- 2. Concave-Convex PDMP (Sutton and Fearnhead, 2023)
- 3. NuZZ (numerical Zig-Zag) (Pagani et al., 2024)

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Slides and codes are available here

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Appendix: Scalability by Subsampling

Construction of ZZ-CV (Zig-Zag with Control Variates).

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3.8 Review: Id Gaussian Mean Reconstruction U' has an alternative form:

$$U'(x) = rac{x}{
ho^2} + rac{1}{\sigma^2}\sum_{i=1}^n (x-y_i) =: rac{1}{n}\sum_{i=1}^n U'_i(x),$$

where

$$U_i'(x)=rac{x}{
ho^2}+rac{n}{\sigma^2}(x-y_i).$$

 \rightarrow We only need one sample y_i to evaluate U'_i .

3.9 Randomized Rate Function

Instead of

$$\lambda_{\mathrm{ZZ}}(x,\sigma) = igg(\sigma U'(x)igg)_+$$

we use

$$\lambda_{ extsf{ZZ-CV}}(x,\sigma) = igg(\sigma U_I'(x)igg)_+, \qquad I \sim \mathrm{U}([n]).$$

Then, the latter is an unbiased estimator of the former:

$$\mathrm{E}_{I \sim \mathrm{U}([n])}igg[\lambda_{\mathrm{ZZ-CV}}(x,\sigma)igg] = \lambda_{\mathrm{ZZ}}(x,\sigma).$$

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3.10 Last Step: Poisson Thinning

Find an invertible upper bound M that satisfies

$$\int_0^t \lambda_{ extsf{ZZ-CV}}(x_s,\sigma_s)\,ds =: m_I(t) \leq oldsymbol{M}(t), \qquad I \sim \mathrm{U}([n]).$$

It is harder to bound λ_{ZZ-CV} , since it is now an estimator (random function).

3.11 Upper Bound M with Control Variates

Preprocessing (once and for all)

I. Find

$$x_*:=rgmin_{x\in\mathbb{R}}U(x)$$

2. Compute

$$U'(x_*) = rac{x_*}{
ho^2} + rac{1}{\sigma^2}\sum_{i=1}^n (x_*-y_i).$$

Then, with a reparameterization of m_i ,

$$m_i(t) \leq M(t) := a + bt,$$

where

$$a = (\sigma U'(x_*))_+ + \|U'\|_{ ext{Lip}} \|x - x_*\|_p, \qquad b := \|U'\|_{ ext{Lip}}.$$

And m_i is redefined as

$$m_i(t) = U'(x_*) + U_i'(x) - U_i'(x_*).$$

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3.12 Subsampling with Control Variates Zig-Zag sampler with the random rate function

$$\lambda_{ extsf{ZZ-CV}}(x,\sigma) = igg(\sigma U_I'(x)igg)_+, \qquad I \sim \mathrm{U}([n]).$$

and the upper bound

$$M(t)=a+bt$$

is called Zig-Zag with Control Variates (Bierkens, Fearnhead, et al., 2019).

3.13 Zig-Zag with Control Variates

I. has O(1) efficiency as the sample size n grows.^I 2. is exact (no bias).



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I. As long as the preprocessing step is properly done.

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3.14 Scalability (1/3)

There are currently two main approaches to scaling up MCMC for large data.

I. Devide-and-conquer

Devide the data into smaller **chunks** and run MCMC on each **chunk**.

2. Subsampling

Use a subsampling estimate of the likelihood, which does not require the entire data.

3.15 Scalability (2/3) by Devide-and-conquer Devide the data into smaller chunks and run MCMC on each chunk.

Unbiased ?	Method	Reference
×	WASP	(Srivastava et al., 2015)
×	Consensus Monte Carlo	(Scott et al., 2016)
\checkmark	Monte Carlo Fusion	(Dai et al., 2019)

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3.16 Scalability (3/3) by Subsampling

Use a subsampling estimate of the likelihood, which does not require the entire data.

Unbiased?	Method	Reference
×	Stochastic Gadient MCMC	(Welling and Teh, 2011)
~	Zig-Zag with Subsampling	(Bierkens, Fearnhead, et al., 2019)
×	Stochastic Gradient PDMP	(Fearnhead et al., 2024)

0 reactions

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0 comments

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